

Research Article





Truncated Akash distribution: properties and applications

Abstract

In this paper, Truncated Akash distribution has been proposed. Its mean and variance have been derived. Nature of cumulative distribution and hazard rate functions have been derived and presented graphically. Its moments including Coefficient of Variation, Skenwness, Kurtosis and Index of dispersion have been derived. Maximum likelihood method of estimation has been used to estimate the parameter of proposed model. It has been applied on three data sets and compares its superiority over one parameter exponential, Lindley, Akash, Ishita and truncated Lindley distribution.

Keywords: Akash distribution, pranav distribution, maximum likelihood estimation, moments, hazard rate

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Introduction

In the recent past decades, life time modeling has been becoming popular in distribution theory, where many statisticians are involved in introducing new models. Some of the life time models are very popular and applied in biological, engineering and agricultural areas, such as Lindley distribution of Lindley, weighted Lindley distribution introduced by Ghitany, Atieh, and Nadarajah, Akash distribution suggested by Shanker, Ishita distribution proposed by Shanker and Shukla, Pranav distribution introduced by Shukla, are some among others and extension of above mentioned distribution has also been becoming popular in different areas of statistics.

Shanker³ proposed Akash distribution convex combination of exponential and gamma distributions which is defined by its pdf and cdf

$$f_1(y;\theta) = \frac{\theta^3}{\theta^2 + 2} (1 + y^2) e^{-\theta y} ; y > 0, \theta > 0$$
 (1.1)

$$F_2(y;\theta) = 1 - \left[1 + \frac{\theta y(\theta y + 2)}{\theta^2 + 2}\right] e^{-\theta y}; y > 0, \theta > 0$$
 (1.2)

The r th moment about origin μ_r of Akash distribution obtained by Shanker is

$$\mu_{r}' = \frac{r! \left\{ \theta^{2} + (r+1)(r+2) \right\}}{\theta^{r} \left(\theta^{2} + 2 \right)} ; r = 1, 2, 3, \dots$$
 (1.3)

Shanker³ has discussed in details about its mathematical and statistical properties, estimation of parameters and applications to model lifetime data from engineering and biomedical engineering.

Truncated type of distribution are more effective for modeling life time data because its limits used as bound either upper or lower or both according to the given data. Truncated normal distribution is proposed by Johnson, Kotz, and Balakrishnan.⁶ It has wide application in economics and statistics. Many researchers have been proposed truncated type of distribution and applied in different areas of statistics, especially in censor data such as truncated Weibull

distribution of Zange and Xie,¹² truncated Lomax distribution of Aryuyuen and Bodhisuwan,⁸ truncated Pareto distribution of Janinetti and Ferraro,⁹ truncated Lindley distribution of Singh, Singh, and Sharma.¹⁰ Truncated version of a continuous distribution can be defined as:

Definition1. Let X be a random variable distributed according to some pdf $g(x;\theta)$ and cdf $G(x;\theta)$, where θ is a parameter vector of X. Let X lies within the interval [a,b], where $-\infty < a \le x \le b < \infty$, then X, conditional on $a \le x \le b$ is distributed as truncated distribution. The pdf of truncated distribution as reported by Singh, Singh, and Sharma¹⁰ defined by:

$$f(x;\theta) = g(x \mid a \le x \le b;\theta) = \frac{g(x;\theta)}{G(b;\theta) - G(a;\theta)}$$
(1.3)

where $f(x;\theta) = g(x;\theta)$ for all $a \le x \le b$ and $f(x;\theta) = 0$ elsewhere.

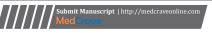
Note that $f(x;\theta)$ in fact is a pdf of X on interval [a,b]. We have

$$f(x;\theta) = \int_{a}^{b} f(x;\theta)dx = \frac{1}{G(b;\theta) - G(a;\theta)} \int_{a}^{b} g(x;\theta)dx$$
$$= \frac{1}{G(b;\theta) - G(a;\theta)} G(b;\theta) - G(a;\theta) = 1$$
(1.4)

The cdf of truncated distribution is given by

$$F(x;\theta) = \int_{a}^{x} f(x;\theta)dx = \frac{G(x;\theta) - G(a;\theta)}{G(b;\theta) - G(a;\theta)}$$
(1.5)

The main objective of this paper is to propose new truncated distribution using Akash distribution, which is called as truncated Akash distribution. It has been divided in seven sections. Introduction about the paper is described in the first section. In the second section, truncated Akash distribution has been derived. Behavior of hazard rate has been presented in third section Statistical properties including its moment have been discussed in the fourth section. Estimation of parameters of the proposed distribution has been discussed in the fifth section. Its application and comparative study with one parameter life time distribution have been illustrated in the section sixth. Finally the conclusion of the paper has been given in the seventh section.





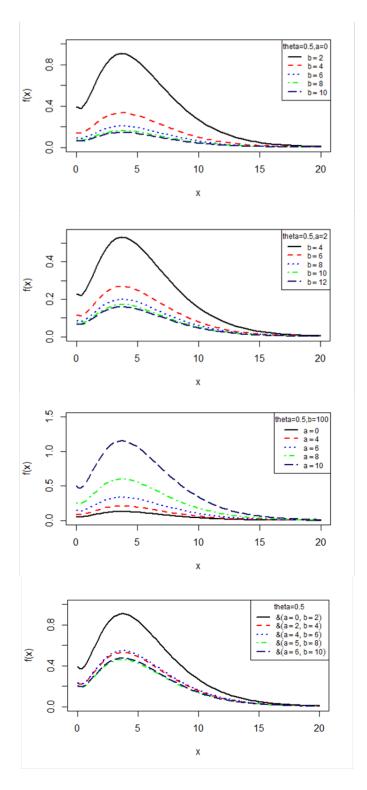


Figure I pdf plots of TAD for varying values of parameters.

Truncated Akash distribution

In this section, pdf and cdf of new truncated distribution is proposed and named Truncated Akash distribution, using (1.3) & (1.4) of definition 1 and from (1.1) & (1.2), which is defined as:

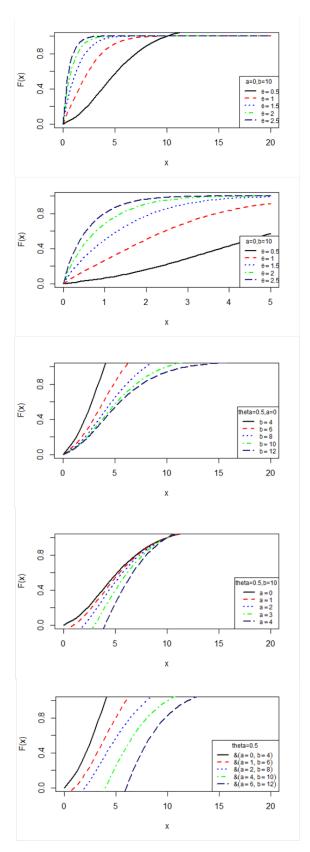


Figure 2 cdf plots of TAD for varying values of parameter.

Definition 2: Let X be random variable which is distributed as Truncated Akash distribution (TAD) with location parameters a, b

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and scale θ and denoted by TAD (a,b,θ) . The pdf and cdf of X are respectively:

$$f(x;\theta) = \frac{\theta^{3}(x^{2}+1)e^{-\theta x}}{a\theta(a\theta+2)e^{-\theta a} - b\theta(b\theta+2)e^{-\theta b} + (\theta^{2}+2)(e^{-\theta a} - e^{-\theta b})}$$
(2.1)

$$F(x;\theta) = \frac{a\theta(a\theta + 2)e^{-\theta a} - x\theta(x\theta + 2)e^{-\theta x} + (\theta^2 + 2)(e^{-\theta a} - e^{-\theta x})}{a\theta(a\theta + 2)e^{-\theta a} - b\theta(b\theta + 2)e^{-\theta b} + (\theta^2 + 2)(e^{-\theta a} - e^{-\theta b})}$$
where $-\infty < a \le x \le b < \infty$, and $\theta > 0$

Performance of pdf and cdf of TAD for varying values of parameters has been illustrated in the figure 1&2 respectively.

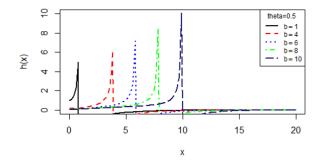
Survival and hazard function

The survival function S(x) and the hazard function h(x) of TAD are defined as

$$S(x) = 1 - F(x) = \frac{x\theta(x\theta + 2)e^{-\theta x} - b\theta(b\theta + 2)e^{-\theta b} + (\theta^2 + 2)(e^{-\theta x} - e^{-\theta b})}{a\theta(a\theta + 2)e^{-\theta a} - b\theta(b\theta + 2)e^{-\theta b} + (\theta^2 + 2)(e^{-\theta a} - e^{-\theta b})}$$

$$h(x) = \frac{f(x)}{S(x)} = \frac{\theta^3(x^2 + 1)e^{-\theta x}}{x\theta(x\theta + 2)e^{-\theta x} - b\theta(b\theta + 2)e^{-\theta b} + (\theta^2 + 2)(e^{-\theta x} - e^{-\theta b})}$$

It is obvious that h(x) is independent from parameter a. Behavior of hazard function of TAD for varying values of parameter is presented in figure 3.



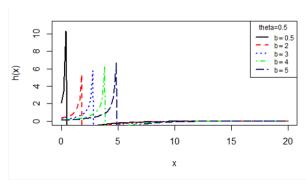


Figure 3 h(x) plots of TAD for varying values of parameter.

Moments and Mathematical Properties

Theorem: Suppose X follows doubly TAD (θ,a,b) . Then the r th moment about origin μ_r' of TAD is

$$\mu_r' = \frac{\theta^2 \left\{ \gamma \left(r+1,\theta b\right) - \gamma \left(r+1,\theta a\right) \right\} + \left\{ \gamma \left(r+2,\theta b\right) - \gamma \left(r+2,\theta a\right) \right\}}{\theta^r \left(a\theta (a\theta + 2)e^{-\theta a} - b\theta (b\theta + 2)e^{-\theta b} + (\theta^2 + 2)(e^{-\theta a} - e^{-\theta b}) \right)}; r = 1,2,3,\dots$$

Proof: Considering

$$K = \left\{ a\theta(a\theta + 2)e^{-\theta a} - b\theta(b\theta + 2)e^{-\theta b} + (\theta^2 + 2)(e^{-\theta a} - e^{-\theta b}) \right\}$$

in (1.6), we have

$$\mu_r' = \frac{\theta^3}{K} \int_a^b x^r (1 + x^2) e^{-\theta x} dx$$
$$= \frac{\theta^3}{K} \left[\int_a^b e^{-\theta x} x^r dx + \int_a^b e^{-\theta x} x^{r+2} dx \right]$$

Taking
$$u = \theta x, x = \frac{u}{\theta}$$

$$= \frac{\theta^{3}}{K} \left[\frac{1}{\theta^{r+1}} \begin{cases} \frac{1}{\theta} e^{-u} x^{r} du - \int_{0}^{\theta a} e^{-u} x^{r} du \\ \frac{1}{\theta^{r+2}} \begin{cases} \frac{\theta b}{\theta} e^{-u} u^{r+2} du - \int_{0}^{\theta a} e^{-u} x^{r+2} du \\ \frac{1}{\theta} e^{-u} u^{r+2} du - \frac{\theta a}{\theta} e^{-u} u^{r+2} du \end{cases} \right]$$

Where $\gamma(\alpha, z) = \int_{0}^{z} e^{-x} x^{\alpha - 1} dx, \alpha > 0, x > 0$ is the lower incomplete

gamma function

$$= \frac{\theta^3}{K} \left[\frac{\gamma(r+1,\theta b) - \gamma(r+1,\theta a)}{\theta^{r+1}} + \frac{\gamma(r+3,\theta b) - \gamma(r+3,\theta a)}{\theta^{r+3}} \right]$$

$$=\frac{1}{K}\left[\frac{\theta^2\left\{\gamma(r+1,\theta b)-\gamma(r+1,\theta a)\right\}+\left\{\gamma(r+3,\theta b)-\gamma(r+3,\theta a)\right\}}{\theta^r}\right]$$

$$= \frac{\theta^2 \left\{ \gamma(r+1,\theta b) - \gamma(r+1,\theta a) \right\} + \left\{ \gamma(r+2,\theta b) - \gamma(r+2,\theta a) \right\}}{\theta^r \left(a\theta(a\theta+2)e^{-\theta a} - b\theta(b\theta+2)e^{-\theta b} + (\theta^2+2)(e^{-\theta a} - e^{-\theta b}) \right)}$$
(4.1)

Now putting r = 1,2 in (4.1), mean and variance can be obtained as

$$\mu_{1}' = \frac{\theta^{2} \{ \gamma(2, \theta b) - \gamma(2, \theta a) \} + \{ \gamma(4, \theta b) - \gamma(4, \theta a) \}}{\theta (a\theta (a\theta + 2)e^{-\theta a} - b\theta (b\theta + 2)e^{-\theta b} + (\theta^{2} + 2)(e^{-\theta a} - e^{-\theta b}))}$$

$$\dot{\mu_{2}} = \frac{\theta^{2} \left\{ \gamma(3,\theta b) - \gamma(3,\theta a) \right\} + \left\{ \gamma(5,\theta b) - \gamma(5,\theta a) \right\}}{\theta^{2} \left(a\theta(a\theta + 2)e^{-\theta a} - b\theta(b\theta + 2)e^{-\theta b} + (\theta^{2} + 2)(e^{-\theta a} - e^{-\theta b}) \right)}$$

Variance
$$\mu_2 = \mu_2 - (\mu_1)^2$$

Similarly rest two moment of origin as well as coefficient of variation, coefficient of skewness, coefficient of kurtosis and Index of dispersion can be obtained, substituting r = 3,4 in the equation (4.1), which are as follows:

$$\mu_{3}' = \frac{\theta^{4} \left\{ \gamma(4,\theta b) - \gamma(4,\theta a) \right\} + \left\{ \gamma(6,\theta b) - \gamma(6,\theta a) \right\}}{\theta^{3} \left(a\theta(a\theta + 2)e^{-\theta a} - b\theta(b\theta + 2)e^{-\theta b} + (\theta^{2} + 2)(e^{-\theta a} - e^{-\theta b}) \right)}$$

$$\mu_{4}^{'} = \frac{\theta^{4} \left\{ \gamma(5,\theta b) - \gamma(5,\theta a) \right\} + \left\{ \gamma(8,\theta b) - \gamma(8,\theta a) \right\}}{\theta^{4} \begin{pmatrix} (a^{3}\theta^{3} + 3a^{2}\theta^{2} + 6a\theta + \theta^{4} + 6)e^{-\theta a} - \\ (b^{3}\theta^{3} + 3b^{2}\theta^{2} + 6b\theta + \theta^{4} + 6)e^{-\theta b} \end{pmatrix}}$$

Coefficient of Variation=
$$\frac{\left(\mu_2^{'}-(\mu_1^{'})^2\right)^{1/2}}{\mu_1^{'}}, \quad \text{Coefficient of Skweness} = \frac{\left(\mu_3^{'}+3\mu_2^{'}\mu_1^{'}-(\mu_1^{'})^2\right)}{\left(\mu_2^{'}-(\mu_1^{'})^2\right)^{3/2}}, \quad \text{Coefficient of Kurtosis} =$$

$$\frac{\left(\dot{\mu_{4}}-4\dot{\mu_{3}}\dot{\mu_{1}}+6\dot{\mu_{2}}(\dot{\mu_{1}})^{2}-3(\dot{\mu_{1}})^{4}\right)}{(\dot{\mu_{2}}-(\dot{\mu_{1}})^{2})^{2}}\ ,$$

Index of dispersion= $\frac{\left(\mu_2' - (\mu_1')^2\right)}{\mu_1'}$, However, they can be easily

Maximum Likelihood Method Estimation

Let $(x_1, x_2, x_3, ..., x_n)$ be a random sample of size n from (1.1). The likelihood function, L of TAD is given by

$$L = \left(\frac{\theta^3}{a\theta(a\theta+2)e^{-\theta a} - b\theta(b\theta+2)e^{-\theta b} + (\theta^2+2)(e^{-\theta a} - e^{-\theta b})}\right)^n \prod_{i=1}^n \left(1 + x_i^2\right) e^{-n\theta \overline{x}}$$
The log likelihood function is thus obtained as

$$\ln L = n \ln \left(\frac{\theta^3}{a\theta(a\theta + 2)e^{-\theta a} - b\theta(b\theta + 2)e^{-\theta b} + (\theta^2 + 2)(e^{-\theta a} - e^{-\theta b})} \right) + \sum_{i=1}^n \ln \left(1 + x_i^2 \right) - n \theta \overline{x}$$

Taking $\hat{a} = \min(x_1, x_2, x_3, ..., x_n)$, $\hat{b} = \max(x_1, x_2, x_3, ..., x_n)$, the

maximum likelihood estimate θ of parameter θ is the solution of the log-likelihood equation $\frac{\partial \log L}{\partial \theta} = 0$. It is obvious that $\frac{\partial \log L}{\partial \theta} = 0$ will not be in closed form and hence some numerical optimization technique can be used e the equation for θ . In this paper the nonlinear method available in R software has been used to find the MLE of the parameter θ .

Applications on Life time data

In this section, TAD has been applied to three datasets using maximum likelihood estimates. Parameter θ is estimated whereas another parameters a, and b are considered as lowest and highest values of data. i. e. $\hat{a} = \min(x_1, x_2, x_3, ..., x_n)$ and $\hat{b} = \max(x_1, x_2, x_3, ..., x_n)$. Goodness of fit has been decided using Akaike information criteria (AIC), Bayesian Information criteria (BIC) and Kolmogorov Simonov test (KS) values respectively, which are calculated for each distribution and also compared with p-value and given in the table 1,2 &3. As we know that best goodness of fit of the distribution can be decided on the basis of minimum value of KS, AIC and BIC.

Table I MLE's, Standard Errors, - 2In L, AIC, BIC, K-S and p-values of the fitted distributions for data set-5

Distributions	ML Estimates	Standard Errors	−2 ln <i>L</i>	AIC	ВІС	K-S	p-value
TAD	$\hat{\theta} = 0.03917$	0.00303	939.13	941.13	942.05	0.153	0.017
TLD	$\hat{\theta} = 0.02199$	0.00273	958.88	960.88	962.31	0.186	0.001
Akash	$\hat{\theta} = 0.04387$	0.00253	950.97	952.97	954.40	0.194	0.001
Ishita	$\hat{\theta} = 0.04390$	0.002533	950.92	9952.92	954.35	0.194	0.001
Lindley	$\hat{\theta} = 0.02886$	0.002038	983.10	985.10	986.54	0.252	0.000
Exponential	$\hat{\theta} = 0.01463$	0.001457	1044.87	1046.87	1048.30	0.336	0.000

Table 2 MLE's, Standard Errors, - 2In L, AIC, BIC, K-S and p-values of the fitted distributions for data set-2

Distributions	ML Estimates	Standard Errors	−2 ln <i>L</i>	AIC	ВІС	K-S	p-value
TAD	$\hat{\theta} = 0.08776$	0.024241	201.96	203.96	205.58	0.112	0.786
TLD	$\hat{\theta} = 0.05392$	0.023917	202.18	204.18	205.61	0.117	0.738
Akash	$\hat{\theta} = 0.09706$	0.01004	240.68	242.68	242.67	0.298	0.005
Ishita	$\hat{\theta} = 0.097328$	80010.0	240.48	242.48	243.48	0.297	0.006
Lindley	$\hat{\theta} = 0.06299$	0.00800	253.98	255.98	256.98	0.365	0.000
Exponential	$\hat{\theta} = 0.032452$	0.00582	274.52	276.52	277.52	0.458	0.000

Table 3 MLE's, Standard Errors, - 2In L, AIC, BIC, K-S and p-values of the fitted distributions for data set-3

Distributions	ML Estimates	Standard Errors	−2 ln <i>L</i>	AIC	ВІС	K-S	p-value
TAD	$\hat{\theta} = 0.70314$	0.18671	110.76	112.76	114.68	0.152	0.079
TLD	$\hat{\theta} = 0.28986$	0.184873	112.19	114.19	115.63	0.157	0.065
Akash	$\hat{\theta} = 0.96472$	0.06460	224.27	226.27	227.27	0.362	0.000
Ishita	$\hat{\theta} = 0.93156$	0.05602	223.14	225.14	226.13	0.330	0.000
Lindley	$\hat{\theta} = 0.65450$	0.05803	238.38	240.38	241.37	0.401	0.000
Exponential	$\hat{\theta} = 0.40794$	0.04911	261.73	263.73	264.73	0.448	0.000

Data Set 1: The data is given by Birnbaum and Saunders¹¹ on the fatigue life of 6061 - T6 aluminum coupons cut parallel to the direction of rolling and oscillated at 18 cycles per second. The data set consists of 100 observations with maximum stress per cycle 31,000 psi. The data (\times 10⁻³) are presented below (after subtracting 65).

5	25	31	32	34	35	38	39	39	40	42	43
43 43	3 44	44	47	47	48	49	49	49	51	54	55
55	55 56	56	56	58	59	59	59	59	59 63		
63	64 64	65	65	65	66	66	66	66	66	67	67
67	68 69	69	69	69	71	71	72	73	73	73	74
74	76 76	77	77	77	77	77	77	79	79	80	81
83	83 84 86	5 86	87	90	91	92	92	92	92 93	94	
97 98	8 98	99	101	103	105	109	136 1	47			

Data Set 2: This data set is the strength data of glass of the aircraft window reported by Fuller, Frieman, Quinn, Quinn, and Carter: 12

18.83 20.8	21.657	23.03	23.23	24.05	24.321	25.5	25.52	25.8	26.69	26.77	
26.78 27.05 39.58 44.045			31.11	33.2	33.73	33.76	33.89	34.76	35.75	35.91 36.98	37.08 37.09

Data Set 3: The following data represent the tensile strength, measured in GPa, of 69 carbon fibers tested under tension at gauge lengths of 20mm, Bader and Priest: 13

	1.312	1.314	1.479	1.552	1.700	1.803	1.861	1.865	1.944	1.958	1.966	1.997	
	2.006	2.021	2.027	2.055	2.063	2.098	2.140	2.179	2.224	2.240	2.253	2.270	2.272
2.274	2.301	2.301	2.359	2.382	2.382	2.426	2.434	2.435	2.478	2.490	2.511	2.514	2.535
2.554	2.566	2.570	2.586	2.629	2.633	2.642	2.648	2.684	2.697	2.726	2.770	2.773	2.800
2.809	2.818	2.821	2.848	2.880	2.954	3.012	3.067	3.084	3.090	3.096	3.128	3.233	3 . 4 3 3
3.585	3.858												

Fitted plots of the considered distributions are presented in Figure 4, 5 and 6, respectively.

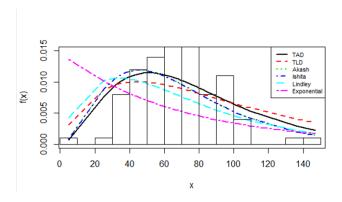


Figure 4 Fitted plots of distributions for the dataset-I.

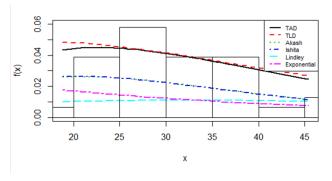


Figure 5 Fitted plot of distributions for data set-2.

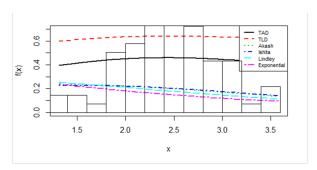


Figure 6 Fitted plot of distribution for data set-3.

Conclusion

In this paper, truncated Akash distribution (TAD) has been proposed. Its statistical properties including survival function and hazard rate have been discussed. Its moments including Coefficient of variation, Skewness, Kurtosis and Index of dispersion have derived. Maximum likelihood method has been used for estimation of its parameter. Goodness of fit of TAD has been discussed with three life time datasets and compared with truncated Lindley, Akash, Ishita, Lindley and exponential distributions. It has been observed that TAD gives good fit over TLD (truncated Lindley Distribution), Akash, Ishita, Lindley and exponential distribution.

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