

# A discrete Pranav distribution and its applications

## Abstract

In the recent decades, the discretization of continuous distribution has been attracting to the attention of researchers because it generates distributions that can be used for strictly discrete data. In this paper, a discrete Pranav distribution, which is a discrete analogue of continuous Pranav distribution, has been carried out. Its important properties including coefficient of variation, skewness, kurtosis and index of dispersion have been obtained and discussed graphically. The method of maximum likelihood estimation has been used for estimating its parameter. The goodness of fit of the proposed distribution have been illustrated using some real count datasets and it was found better fit as compared to other one parameter discrete distributions.

**Keywords:** Pranav distribution, discretization, moment generating function, moments, estimation, goodness of fit

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## Introduction

In recent past years, the use of discrete analogue of a continuous distribution avoids the use of a continuous distribution in the case of strictly available discrete data.

In many cases, it is not easy to get samples from continuous distributions. The observed values, in the most of cases, are collected actually discrete in nature for the reason that they are measured to only finite number of decimal places and cannot completely presents all points in a continuum. According to Lai,<sup>1</sup> discretization of a continuous lifetime model is an appealing approach to derive a discrete lifetime model corresponding to the continuous one. Therefore, it is reasonable and convenient to model the situation by an appropriate discrete distribution generated from the underlying continuous distribution preserving one or more important characteristics including probability density function (pdf), mean residual life function etc. and important statistical properties of the distribution.

In Statistics literature, different researchers have used different methods of discretization to propose a discrete type of distribution analogues of continuous distribution. In this study, one of the discretization methods has been used to find discrete analogue of continuous Pranav distribution introduced by Shukla.<sup>2</sup> Infinite series method has been used to find pmf of Pranav distribution, appropriate definition is given in the next paragraph. This method was firstly used by Good<sup>3</sup> who proposed the discrete Good distribution to model for the frequencies of species. It was given as follows.

A random variable  $Y$  is said to have a discrete Good distribution if its pmf can be expressed as

$$P(Y = y) = \frac{\alpha^y y^\beta}{\sum_{j=0}^{\infty} \alpha^j j^\beta}; y = 0, 1, 2, \dots \quad (1.1)$$

where  $\beta \in R$  and  $\alpha \in (0, 1)$ .

The method of infinite series is formulated by the definition which is given as below:

**Definition:** Let  $X$  be a continuous random variable having pdf  $f_X(x)$  with support on  $R$ . Then the corresponding discrete random variable  $Y$  has pmf given by

$$P(Y = y) = P(y; \theta) = \frac{f_X(y; \theta)}{\sum_{j=-\infty}^{\infty} f_X(j; \theta)}; y \in Z, \quad (1.2)$$

where  $\theta$  may be the vector of parameters indexing the distribution of  $X$ .

This method has been used by many researchers to derive discretization of continuous distribution, such as, Kulasekara & Tonkyn,<sup>4</sup> Doray & Luong,<sup>5</sup> Sato et al.,<sup>6</sup> Nekoukhou et al.,<sup>7</sup> are some among others, who proposed a version of the method when the continuous random variable of interest is defined on  $R_+$ . Thus, if the random variable  $X$  is defined on  $R_+$ , the pmf of  $Y$  can be defined as

$$P(Y = y) = P(y; \theta) = \frac{f_X(y; \theta)}{\sum_{j=0}^{\infty} f_X(j; \theta)}; y \in Z_+ \quad (1.3)$$

Josmar et al.,<sup>8</sup> using infinite series method has derived a discrete Shanker distribution (DSD) with parameter  $\theta > 0$  and having pmf

$$P_1(y; \theta) = \frac{(e^\theta - 1)^2 (\theta + y) e^{-\theta(y+1)}}{1 + (e^\theta - 1)\theta}; y = 0, 1, 2, \dots \quad (1.4)$$

They have discussed its various statistical properties including its applications to model count datasets in their paper. Which is a discrete analogue of continuous Shanker distribution introduced by Shanker<sup>9</sup> having pdf

$$f_1(x; \theta) = \frac{\theta^2}{\theta^2 + 1} (\theta + x) e^{-\theta x}; x > 0, \theta > 0 \quad (1.5)$$

Using same method of discretization, the pmf of discrete Lindley distribution (DLD) proposed by Berhane & Shanker<sup>10</sup> is given by

$$P_2(y; \theta) = \frac{(e^\theta - 1)^2}{e^{2\theta}} (1 + y) e^{-\theta y}; y = 0, 1, 2, \dots \quad (1.6)$$

where the parameter  $\theta > 0$ .

They have discussed its important statistical properties including estimation of parameter of DLD and applied on some count datasets from engineering and biology in their paper. They showed its superiority over other discrete one parameter distributions such as Poisson Lindley distribution (PLD) proposed by Shankar,<sup>11</sup> Poisson Akash distribution (PAD) proposed by Shanker,<sup>12</sup> and DSD proposed by Josmer et al.<sup>8</sup> As mentioned above, DLD is a discrete analogue of continuous Lindley distribution introduced by Lindley<sup>13</sup> having pdf

$$f_2(x; \theta) = \frac{\theta^2}{\theta + 1} (1 + x) e^{-\theta x}; x > 0, \theta > 0. \quad (1.7)$$

Recently, Berhane & Shanker,<sup>14</sup> proposed a discrete Akash distribution (DAD) using infinite series method, the pmf of a discrete Akash distribution is given as

$$P_3(y;\theta) = \frac{(e^\theta - 1)^3}{e^\theta(e^{2\theta} - e^\theta + 2)}(1 + y^2)e^{-\theta y}; y = 0, 1, 2, \dots \quad (1.8)$$

They have discussed its important statistical properties including estimation of method and applied on some count datasets, and showed its superiority over DSD, DLD, PLD and PAD in their paper. Which is the discrete analogue of a continuous Akash distribution introduced by Shanker,<sup>15</sup> its pdf is given as:

$$f_3(x;\theta) = \frac{\theta^3}{\theta^2 + 2}(1 + x^2)e^{-\theta x}; x > 0, \theta > 0 \quad (1.9)$$

Shanker<sup>12</sup> proposed PAD, a Poisson mixture of Akash distribution, having pmf

$$P_4(x;\theta) = \frac{\theta^3}{\theta^2 + 2} \cdot \frac{x^2 + 3x + (\theta^2 + 2\theta + 3)}{(\theta + 1)^{x+3}}; x = 0, 1, 2, \dots, \theta > 0 \quad (1.10)$$

He has discussed its important statistical properties including estimation of parameter along with applications of PAD in his paper. PAD was applied to count datasets and showed its superiority with PLD and other distribution of one parameter.

The PLD is a Poisson mixture of Lindley distribution introduced by Sankaran<sup>11</sup> and is defined by its pmf

$$P_5(x,\theta) = \frac{\theta^2(x + \theta + 2)}{(\theta + 1)^{x+3}}; x = 0, 1, 2, \dots, \theta > 0 \quad (1.11)$$

The main objective of this paper is to propose a discretization of Pranav distribution for the reason being that it was observed, Pranav distribution gives better fit than one parameter continuous Akash, Shaker, Sujatha, Lindley and Exponential distributions. Keeping this view in mind, it is hoped that it would be better than discrete Akash, discrete Shanker and discrete Lindley distributions and other one parameter discrete distributions.

In this paper, Discrete Pranav distribution has been proposed in the second section. It's important properties including coefficients of variation, skewness, kurtosis and index of dispersion have been derived and discussed graphically in the third section. The method of maximum likelihood estimation has been discussed in fourth section. The applications of proposed distribution have been illustrated on two count datasets in fifth section. Conclusions on the basis of result have been discussed in the last section.

### A discrete Pranav distribution

The pdf and the cdf of a random variable  $X$  having Pranav distribution proposed by Shukla<sup>2</sup> are given by

$$f(x;\theta) = \frac{\theta^4}{\theta^4 + 6}(\theta + x^3)e^{-\theta x}; x > 0, \theta > 0 \quad (2.1)$$

$$F(x;\theta) = 1 - \left[ 1 + \frac{\theta x(\theta^2 x^2 + 3\theta x + 6)}{\theta^4 + 6} \right] e^{-\theta x}; x > 0, \theta > 0 \quad (2.2)$$

Shukla<sup>2</sup> has discussed its various statistical properties including moments based measures, hazard rate function, and other important properties along with Bonferroni and Lorenz curves and stress-strength reliability. Pranav distribution applied for modeling lifetime data from biomedical sciences and engineering and explained its

superiority over Akash, Shanker, Ishita, Sujatha, and exponential distributions.

Using the above definition, the pmf of the discrete random variable  $Y$ , corresponding to a continuous random variable  $X$  follows Pranav distribution (2.1) with parameter  $\theta > 0$ , can be obtained as

$$P(y;\theta) = \frac{(e^\theta - 1)^4}{e^\theta(\theta(e^\theta - 1)^3 + e^{2\theta} + 4e^\theta + 1)}(\theta + y^3)e^{-\theta y}; y = 0, 1, 2, \dots \quad (2.3)$$

We would call this distribution, a discrete Pranav distribution (DPD). The nature and behavior of DPD for varying values of its parameter  $\theta$  has been shown graphically in Figure 1. From the figure it was observed that pmf of DPD is increasing as increased values of  $\theta$  (Figure 1).

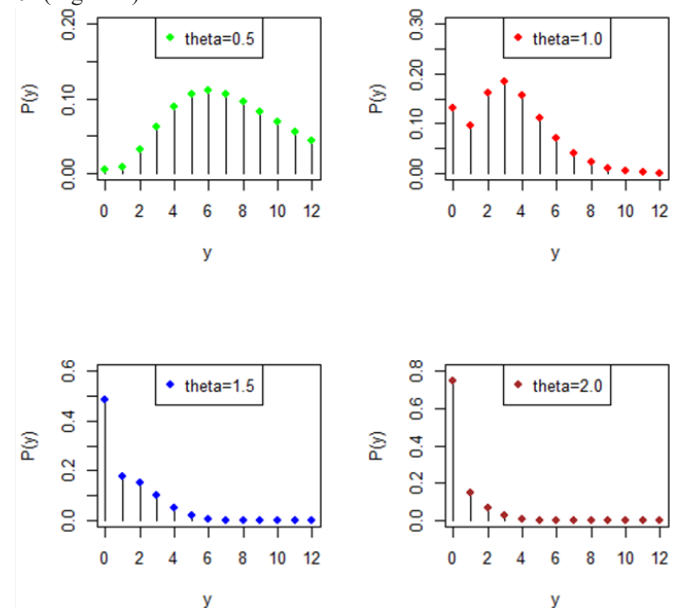


Figure 1 The pmf plot of DPD for varying values of the parameter  $\theta$ .

The survival function,  $S(y;\theta)$  and the cumulative distribution function (cdf),  $F(y;\theta)$  of DPD can be obtained as

$$S_1(y;\theta) = \left[ 1 + \frac{y^3(e^\theta - 1)^3 + 3e^\theta y^2(e^\theta - 1)^2 + 3ye^\theta(e^{2\theta} - 1)}{\theta(e^\theta - 1)^3 + (e^{2\theta} + 4e^\theta + 1)} \right] e^{-\theta(y+1)}; y = 0, 1, 2, \dots, \theta > 0 \quad (2.5)$$

$$F_2(y;\theta) = 1 - \left[ 1 + \frac{y^3(e^\theta - 1)^3 + 3e^\theta y^2(e^\theta - 1)^2 + 3ye^\theta(e^{2\theta} - 1)}{\theta(e^\theta - 1)^3 + (e^{2\theta} + 4e^\theta + 1)} \right] e^{-\theta(y+1)}; y = 0, 1, 2, \dots, \theta > 0$$

cdf graphs of DPD are presented in Figure 2.

Since  $\frac{P(y+1;\theta)}{P(y;\theta)} = \left[ 1 + \frac{3y^2 + 3y + 1}{\theta + y^2} \right] e^{-\theta}$  is a decreasing function of  $y \geq 3$ ,  $P(y;\theta)$  is log-concave and therefore, the DPD has an

increasing hazard rate. Further,  $[P(y;\theta)]^2 \geq P(y-1;\theta) \cdot P(y+1;\theta)$  for  $y \geq 3$ , which implies unimodality, by theorem 3 of Keilson & Gerber.<sup>16</sup> The detailed about interrelationship between log-concavity, unimodality and increasing hazard rate of discrete distributions can be shown in Grandell.<sup>17</sup>

### Mean variance and statistical constants

The probability generating function and the moment generating function (mgf) of DPD can be obtained as

$$G(t) = \frac{(e^\theta - 1)^4}{(\theta(e^\theta - 1)^3 + e^{2\theta} + 4e^\theta + 1)} \left[ \frac{\theta(e^\theta - t)^3 + t(e^{2\theta} + 4te^\theta + t^2)}{(e^\theta - t)^4} \right], \text{ for } t \neq e^\theta$$

and

$$M(t) = \frac{(e^\theta - 1)^4}{\theta(e^\theta - 1)^3 + e^{2\theta} + 4e^\theta + 1} \left[ \frac{\theta(e^\theta - e^t)^3 + e^t(e^{2\theta} + 4e^t e^\theta + e^{2t})}{(e^\theta - e^t)^4} \right], \text{ for } t \neq \theta$$

$$\mu_3 = \frac{\left( \theta^2(\theta+1)e^{10\theta} - (8\theta^3 - 47\theta^2 + \theta)e^{9\theta} + (27\theta^3 - 96\theta^2 + 15\theta)e^{8\theta} - (48\theta^3 + 342\theta^2 + 60\theta - 8)e^{7\theta} + (42\theta^3 + 1320\theta^2 + 300\theta + 32)e^{6\theta} - (1536\theta^2 + 390\theta - 288)e^{5\theta} - (42\theta^3 - 438\theta^2 + 78\theta - 536)e^{4\theta} + (48\theta^3 + 486\theta^2 + 204\theta + 536)e^{3\theta} - (27\theta^3 + 393\theta^2 - 132\theta - 288)e^{2\theta} + (8\theta^3 + 65\theta^2 - 150\theta + 32)e^\theta - \theta^3 + 10\theta^2 - 17\theta + 8 \right)}{(\theta(e^\theta - 1)^3 + e^{2\theta} + 4e^\theta + 1)^3 (e^\theta - 1)^3}$$

$$\mu_4 = \frac{\left( \theta^3(\theta+1)e^{14\theta} - (5\theta^4 - 106\theta^3 + \theta^2)e^{13\theta} - (17\theta^4 + 41\theta^3 - 111\theta^2 + \theta)e^{12\theta} + (230\theta^4 - 3184\theta^3 + 606\theta^2 + 92\theta)e^{11\theta} - (979\theta^4 - 12857\theta^3 + 578\theta^2 + 1136\theta)e^{10\theta} + (2453\theta^4 - 19414\theta^3 - 6363\theta^2 + 5620\theta + 208)e^{9\theta} - (4125\theta^4 - 2391\theta^3 - 6813\theta^2 - 165\theta + 2232)e^{8\theta} + (4884\theta^4 + 34800\theta^3 + 22116\theta^2 - 13416\theta + 8400)e^{7\theta} - (4125\theta^4 + 55749\theta^3 + 49788\theta^2 + 17856\theta - 20544)e^{6\theta} + (2453\theta^4 + 41054\theta^3 + 33561\theta^2 + 26664\theta + 30528)e^{5\theta} - (979\theta^4 + 13987\theta^3 + 3047\theta^2 - 6387\theta - 20544)e^{4\theta} + (230\theta^4 + 128\theta^3 - 3522\theta^2 - 5236\theta + 8400)e^{3\theta} - (17\theta^4 - 1291\theta^3 + 642\theta^2 + 2864\theta - 2232)e^{2\theta} - (5\theta^4 + 242\theta^3 - 707\theta^2 + 668\theta - 208)e^\theta + \theta^4 - 11\theta^3 + 27\theta^2 - 25\theta + 8 \right)}{(\theta(e^\theta - 1)^3 + e^{2\theta} + 4e^\theta + 1)^4 (e^\theta - 1)^4}$$

The coefficient of variation (C.V), coefficient of skewness ( $\sqrt{\beta_1}$ ), coefficient of kurtosis ( $\beta_2$ ) and index of dispersion ( $\gamma$ ) of DPD can be obtained using the relationships below

$$C.V = \frac{\sigma}{\mu_1'} \quad \sqrt{\beta_1} = \frac{\mu_3}{(\mu_2')^{3/2}} \quad \beta_2 = \frac{\mu_4}{\mu_2'^2} \quad \gamma = \frac{\sigma^2}{\mu_1'}$$

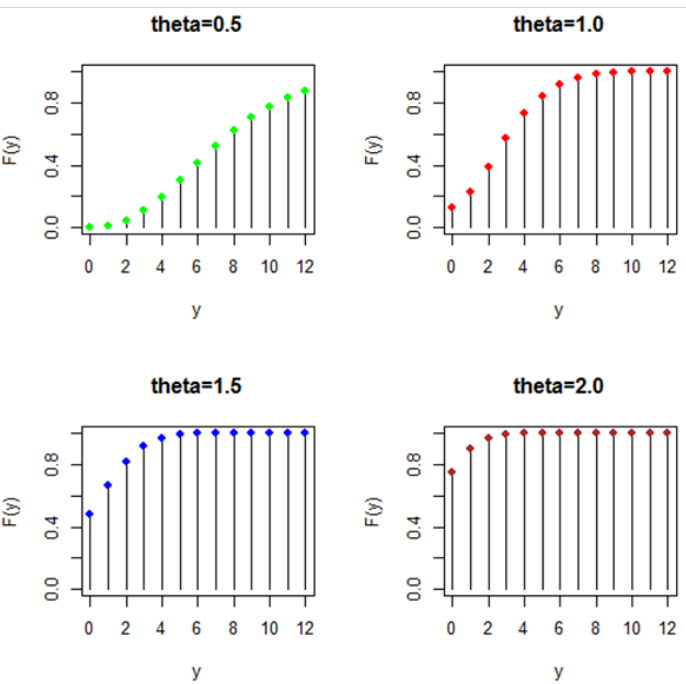


Figure 2 The cdf plot of DPD for varying values of the parameter  $\theta$ .

The first four moments about origin of DPD can thus be obtained as

$$\mu_1' = \frac{\theta(e^\theta - 1)^3 + e^{3\theta} + 11e^{2\theta} + 11e^\theta + 1}{(e^\theta - 1)(\theta(e^\theta - 1)^3 + e^{2\theta} + 4e^\theta + 1)}$$

$$\mu_2' = \frac{\theta e^{4\theta} + e^{4\theta} - 2\theta e^{3\theta} + 26e^{3\theta} + 66e^{2\theta} + 2\theta e^\theta + 26e^\theta - \theta + 1}{(e^\theta - 1)^2(\theta(e^\theta - 1)^3 + e^{2\theta} + 4e^\theta + 1)}$$

$$\mu_3' = \frac{\theta e^{5\theta} + e^{5\theta} + \theta e^{4\theta} + 57e^{4\theta} - 8\theta e^{3\theta} + 302e^{3\theta} + 8\theta e^{2\theta} + 302e^{2\theta} - \theta e^\theta + 57e^\theta - \theta + 1}{(e^\theta - 1)^3(\theta(e^\theta - 1)^3 + e^{2\theta} + 4e^\theta + 1)}$$

$$\mu_4' = \frac{\theta e^{6\theta} + e^{6\theta} + 8\theta e^{5\theta} + 120e^{5\theta} - 19\theta e^{4\theta} + 1191e^{4\theta} + 2416e^{3\theta} + 19\theta e^{2\theta} + 1191e^{2\theta} - 8\theta e^\theta + 120e^\theta - \theta + 1}{(e^\theta - 1)^4(\theta(e^\theta - 1)^3 + e^{2\theta} + 4e^\theta + 1)}$$

Using the relationship  $\mu_r = E(Y - \mu_1')^r = \sum_{k=0}^r \binom{r}{k} \mu_k' (-\mu_1')^{r-k}$  between central moments and moments about origin, the central moments of DPD are derived as

$$\mu_2 = \frac{e^\theta \left( (\theta+1)\theta e^{6\theta} + (22-6\theta)\theta e^{5\theta} + (15\theta^2 - 23\theta + 8)e^{4\theta} - (20\theta^2 + 64\theta - 28)e^{3\theta} + (15\theta^2 + 95\theta + 72)e^{2\theta} - (6\theta^2 + 22\theta - 28)e^\theta + \theta^2 - 9\theta + 8 \right)}{(e^\theta - 1)^2(\theta(e^\theta - 1)^3 + e^{2\theta} + 4e^\theta + 1)^2}$$

Table 1 exhibits the nature and behavior of coefficient of variation (C.V), coefficient of skewness, coefficient of kurtosis and index of dispersion (ID) for varying values of the parameter  $\theta$  (Figure 3).

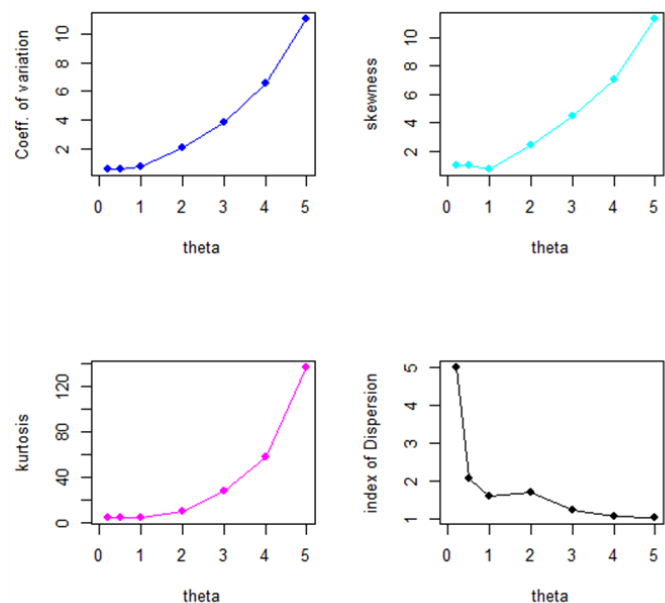


Figure 3 The plot of measures of descriptive statistics of DPD for varying values of the parameter  $\theta$ .

It is clear from Table 1 that the mean, variance, and index of dispersion of DPD are decreasing as increased values of the parameter  $\theta$ , whereas coefficient of variation, coefficient of skewness and coefficient of kurtosis of DPD are increasing as increased values of parameter  $\theta$ .  $\sigma^2 > \mu$ , indicates that DPD can be a suitable model for over-dispersed data.

**Table 1** Values of descriptive statistics of DPD for varying values of  $\theta$

$\theta$	Values of descriptive statistics					
	Mean	Variance	C.V	Skewness	Kurtosis	ID
0.5	7.915	16.3833	0.5114	0.95	4.4328	2.0699
1	3.2844	5.288	0.7001	0.6689	3.6405	1.61
1.5	1.1918	2.2352	1.2545	1.3138	4.5122	1.8755
2	0.4147	0.7018	2.0197	2.4253	9.6591	1.692
2.5	0.1712	0.2418	2.8725	3.4854	17.7216	1.4125
3	0.0825	0.1013	3.8558	4.479	27.2341	1.2272
3.5	0.0438	0.0492	5.0675	5.5959	39.358	1.1239
4	0.0245	0.0261	6.6069	7.0201	57.6163	1.0682

### Maximum likelihood estimation

The likelihood function,  $L$  of (2.3) is given by

$$L = \left( \frac{(e^\theta - 1)^4}{e^\theta (\theta (e^\theta - 1)^3 + e^{2\theta} + 4e^\theta + 1)} \right)^n \prod_{i=1}^n (\theta + y_i^3) e^{-\theta y_i}$$

**Table 2** Distribution of mistakes in copying groups of random digits

No. of error per group	Observed	Expected frequency					
	Frequency	PLD	PAD	DLD	DSD	DAD	DPD
0	35	33.1	33.5	31	31.7	33.2	36
1	11	15.3	14.7	17.4	16.9	14.2	10.6
2	8	6.7	6.6	7.4	7.2	7.6	7.1
3	4	2.9	3	2.8	2.7	3.3	3.9
4	2	2	2.2	1.4	1.5	1.7	2.4
Total	60	60	60	60	60	60	60
	$\hat{\theta}$	1.7434	2.078	1.2678	1.2276	1.5404	1.689
	$\chi^2$	1.8141	1.4185	3.3667	2.9963	1.0398	0.1712
	d.f.	1	2	1	1	2	2
	p-value	0.178	0.492	0.066	0.0837	0.595	0.919

**Table 3** Observed and expected frequencies for distribution of *Pyrausta nublialis* in 1937

No. of insects	Observed	Expected frequency					
	Frequency	PLD	PAD	DLD	DSD	DAD	DPD
0	33	31.5	32	29.6	30.3	31.6	34.4
1	12	14.2	13.6	16.2	15.6	13.2	9.8
2	6	6.1	6	6.6	6.4	6.9	6.4
3	3	2.5	2.6	2.4	2.4	2.9	3.4
4	1	1	1.1	0.8	0.8	1	1.4
5	1	0.7	0.7	0.4	0.5	0.4	0.6
Total	56	56	56	56	56	56	56
	$\hat{\theta}$	1.8115	2.1446	1.2993	1.2535	1.5686	1.7122
	$\chi^2$	0.4598	0.2541	1.5422	1.1516	0.1747	0.6055
	d.f.	1	1	1	1	1	2
	p-value	0.498	0.614	0.215	0.283	0.676	0.739

and its log likelihood function is

$$\ln L = n \ln \left( \frac{(e^\theta - 1)^4}{e^\theta (\theta (e^\theta - 1)^3 + e^{2\theta} + 4e^\theta + 1)} \right) + \sum_{i=1}^n \ln(\theta + y_i^3) - n\theta \bar{y}$$

$$\ln L = 4n \ln(e^\theta - 1) - \theta - \ln(\theta (e^\theta - 1)^3 + e^{2\theta} + 4e^\theta + 1) + \sum_{i=1}^n \ln(\theta + y_i^3) - n\theta \bar{y}$$

Differentiating above equation with respect  $\theta$ , we have

$$\frac{\partial \ln L}{\partial \theta} = \frac{4ne^\theta}{e^\theta - 1} - 1 - \frac{(e^\theta - 1)^3 + 3\theta e^\theta (e^\theta - 1)^2 + 2e^{2\theta} + 4e^\theta}{\theta(e^\theta - 1)^3 + e^{2\theta} + 4e^\theta + 1} + \sum_{i=1}^n \frac{1}{\theta + y_i^3} - n\bar{y} = 0$$

Above equation can be simplified to solve the value of parameter. In this paper, R-software is used to estimate for value of  $\theta$ .

### Application and goodness of fit

In this section, the goodness of fit of the DPD has been discussed with two count datasets. The dataset in Table 2 has been taken from Kemp & Kemp<sup>18</sup> and dataset in Table 3 has been taken from Beall,<sup>19</sup> detailed about the datasets can be shown in their paper. The proposed model is compared with DSD, DLD, PLD, PAD and DAD (Figure 4&5).<sup>20</sup>

It is observed from (Table 2&3) that DPD gives much closer fit than other considered discrete distributions. Hence, it can be considered an important discrete distribution for biological data over other one parameter discrete distributions.

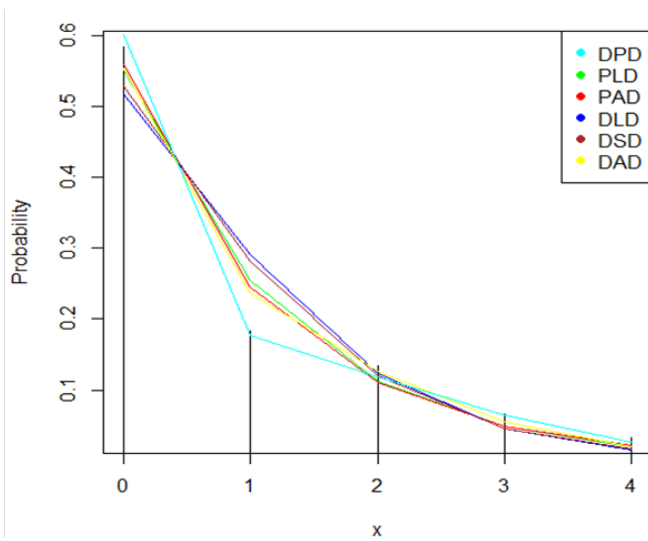


Figure 4 Fitted plot of distributions on first data set.

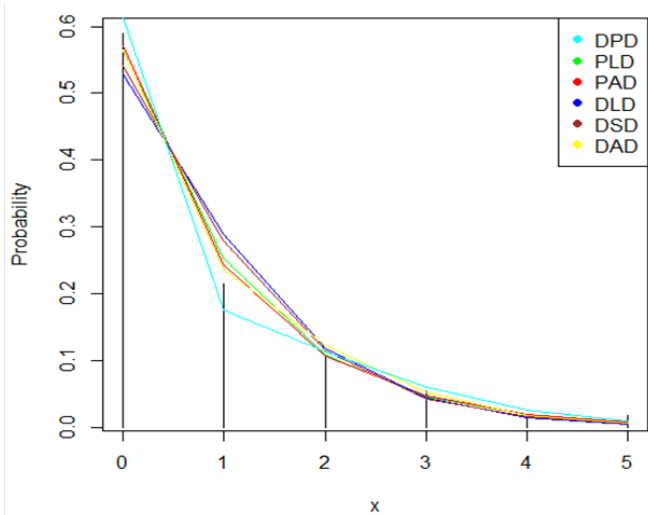


Figure 5 Fitted plot of distributions on second data set.

### Conclusion

In this paper, a discrete Pranav distribution (DPD) has been proposed. Its moment generating function, moments and moments based measures including statistical constants have been derived and their nature and behavior has been discussed numerically and graphically. The method of maximum likelihood estimation has been discussed for estimating its parameter. The goodness of fit of DPD has been explained using two real count datasets. The DPD gives better fit as compared to PLD, PAD, DLD, DSD and DAD in the presented datasets.

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### Conflicts of interest

Author declares that there is no conflict of interest.

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